



## Recognition of micro-scale deformation structures in glacial sediments – pattern perception, observer bias and the influence of experience

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It is a scientist's mission to try to remain unbiased. However, certain factors play a role in scientific analyses that are not controlled by conscious thought. These factors are potentially very important in areas of science where interpretations are based on a scientist's ability to identify patterns or structures. One such area is the micromorphology of glacial sediments. In this paper we investigate the role of an analyst's experience in relation to pattern perception with specific reference to turbate microstructures in glacial diamictons. An experiment was conducted on 52 participants, which demonstrated that, as may be expected, more experienced (glacial) micromorphologists tend to exhibit a higher sensitivity-to-signal, but that complete novices, if given clear instructions, can reach levels of sensitivity similar to those of experts. It also showed, perhaps more surprisingly, that response bias does not decrease with experience. We discuss psychological factors, such as the drive for success and consistency, that may have contributed to these results and investigate their possible implications in the micromorphological analysis and interpretation of glacial sediments.

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The subglacial environment still presents glaciologists and glacial sedimentologists with a great challenge. Owing to its inaccessibility, the understanding of processes operating at the ice-bed interface and within the glacier substrate is still incomplete, and – as a consequence – how, and to what extent, such processes control ice dynamics can sometimes only be guessed at. Given the lack of *in situ* observations and measurements from modern glacial environments, palaeoglaciologists have tried to draw inferences from past glacial sedimentary records. A technique that has been very successful in the interpretation of such palaeorecords is micromorphology, or the study of thin sections (cf. Hiemstra 2006). Micromorphology has proved particularly successful in distinguishing between different glacial diamictons, particularly between those that, when viewed with the naked eye, appear identical. When viewed through the microscope, however, many of these apparently massive diamictons show micro-scale features that are indicative or even diagnostic of certain processes, thus allowing interpretations to be made.

Within this approach it is imperative that observations are accurate and that the micro-scale features referred to are identified for what they are or represent. One micro-structure that is generally taken as compelling evidence for (subglacial) deformation of glacial diamictons is the turbate micro-structure. It consists of a circular (or partly circular) arrangement of often silt- or sand-sized grains. Long axes of these fine grains may be aligned with the overall circular shape, while the

structure may or may not contain a coarser 'core' grain in its centre. Turbate structures, often hundreds of micrometres, and occasionally up to several millimetres, in diameter, are thought to reflect micro-scale rotational movements of rigid elements within a deforming plastic or viscous sediment that re-align small grains in their vicinity (van der Meer 1993, 1997). Studies that have referred to turbate micro-structures in this context include but are not limited to Hiemstra & Rijdsdijk (2003), Carr *et al.* (2006), Hart (2006), Phillips (2006), Larsen *et al.* (2007), Kilfeather *et al.* (2010), Menzies *et al.* (2010) and Menzies & Ellwanger (2011).

There is an inherent human element associated with micromorphological analysis. Although there have been attempts to quantify or automate certain aspects (see Stroeven *et al.* 2005; Zaniewski & van der Meer, 2005; Larsen *et al.* 2007; Phillips *et al.* 2011; Reinardy *et al.* 2011), informed judgments by an observer or observers remain the basis of the method. It is therefore crucial to consider to what extent these judgments may be skewed by limitations of human perception, decision-making or categorization, and to what extent such limitations may be mitigated by experience.

Performance in detection tasks can be evaluated with respect to the observer's (or analyst's) sensitivity-to-signal (i.e. the ability to discriminate the presence versus the absence of a signal) and specificity or response bias (i.e. the general tendency to make a positive response when asked to carry out a survey task). In the case of human observers, both of these aspects could, in principle, change with experience. For

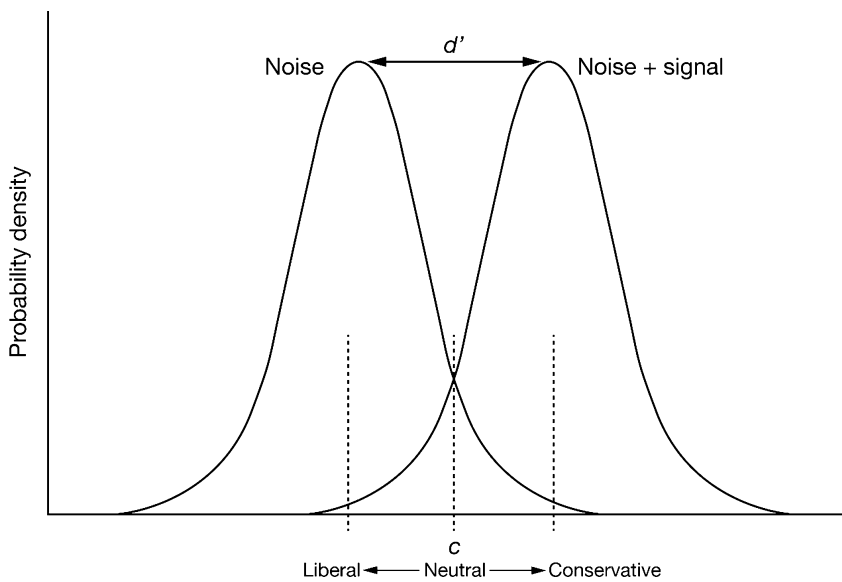


Fig. 1. Probability density distributions of responses in detection experiments. The curve on the left represents 'noise alone' and the curve on the right represents 'signal plus noise'. Indicated are the graphical representation of sensitivity-to-signal  $d'$  and three possible forms of bias  $c$  (liberal, neutral and conservative). Note that a liberal response bias includes most of the signal (hits) but also a large amount of noise, and, conversely, a conservative bias includes little noise but misses a large proportion of signal. See text for further explanation.

example, sensitivity may increase with experience, or experience may lead to an increasingly neutral response bias (see below). In the present study, we test human observers' ability to detect turbate micro-structures (the signal, or stimulus) in photomicrographs (the trials) of glacial diamictos and evaluate how this ability is influenced by experience. In an effort to fully characterize detection performance, we go beyond simply investigating detection accuracy by characterizing observers' performance in terms of sensitivity and response bias.

### Signal detection theory and glacial micromorphology

Responses in (visual) detection tasks can be classified as 'hits' (correct detections of a signal), 'misses' (failures to detect a signal), 'false alarms' (erroneous detections, i.e. positive responses in the absence of a signal) and 'correct rejections' (negative responses in the absence of a signal). In detection experiments, as routinely used in the field of psychology, the total proportion of trials containing the signal is known, and thus detection performance can be fully described with only one of the response proportions from trials with and without a signal (usually, hits and false alarms). Deriving indices of sensitivity and response bias from hits and false alarms requires assumptions about relevant perceptual and response processes. Signal Detection Theory (SDT; Green & Swets 1966) represents a popular framework for evaluating performance in detection tasks in terms of sensitivity and response bias. Although it has not been without critics (see Mueller & Weidemann 2008, for a review), it is well suited to the present study. In a

nutshell, SDT assumes that perception is noisy, such that multiple presentations of identical sensory input are likely to generate (slightly) different percepts (variance in sensory inputs further increases perceptual variance). Thus trials containing a signal (plus noise) give rise to a distribution of percepts that can overlap with the distribution of percepts from trials without a signal (noise alone; Fig. 1). In its simplest form, SDT assumes that percepts for trials with and without a signal are distributed normally with equal variance along a relevant dimension in perceptual space. A response criterion ( $c$  in Fig. 1) then maps percepts to responses and the placement of this response criterion determines the specificity, or response bias. A liberal response criterion leads to many positive responses and therefore ensures a high hit rate (at the cost of a high false alarm rate). In this scenario one would expect a larger number of signals (in our experiment: turbate micro-structures) to be identified, both correctly and incorrectly. A conservative response criterion, on the other hand, leads to few positive responses, thus reducing the number of false alarms, but also reducing the number of hits. In other words, the total number of identified turbate micro-structures would be low, minimizing the mistaken classification of samples that do not contain such micro-structures. However, in this case, the analyst also runs the risk of dismissing crucial evidence when it comes to the interpretation of the sediment in question and is likely to miss turbate micro-structures in the sample. How hits and false alarms trade off across different placements of the response criterion is determined by the overlap between the sensory distributions for trials with and without a signal. In short, the observer's sensitivity increases as the overlap between these distributions decreases (Fig. 1; Green & Swets 1966).

In the context of SDT, several measures have been proposed to quantify sensitivity and response bias. For the purpose of the present research, we have adopted the relatively simple measures of  $d'$  (sensitivity) and  $c$  (for 'criterion', which is an index of response bias). The exact theoretical background of the parameterization is beyond the scope of this paper. Here it suffices to say that the calculation of  $d'$  consists of taking the inverse normal transformation of the hit and false alarm rates and subtracting the latter from the former, as shown in Equation (1):

$$d' = Z(H) - Z(F), \quad (1)$$

where  $Z(x)$  denotes the inverse normal ('Z') transformation and  $H$  and  $F$  denote the hit and false alarm rates respectively.

This measure gives the distance between the two perceptual distributions (for trials with and without signal; see Fig. 1) in units of their standard deviation. In other words, a  $d'$  score of 1 implies that the mean of the distribution of percepts corresponding to observations containing a signal (e.g. turbate micro-structures) is one standard deviation larger than that corresponding to observations not containing a signal. The greater the distance between these distributions ( $d'$ ), the smaller their overlap and hence, all else being equal, the smaller the proportion of misclassified samples. The exact proportions of hits and false alarms are then determined by the placement of the decision criterion  $c$  (Fig. 1). Effectively,  $c$  can be thought of as a threshold on the strength of the percept such that every time a percept exceeds this threshold (loosely corresponding to it 'looking enough' like a signal), the corresponding sample is classified as containing the signal. The criterion  $c$  is calculated by multiplying the sum of the Z-transformed hit and false alarm rates by  $-0.5$ , as shown in Equation (2):

$$c = -0.5\{Z(H) + Z(F)\}, \quad (2)$$

where  $Z(x)$  denotes the inverse normal ('Z') transformation and  $H$  and  $F$  denote the hit and false alarm rates respectively.

The criterion  $c$  measures the response bias relative to an ideal observer: a neutral value of 0 implies that equal priority is given to the minimization of misses and false alarms (Fig. 1). Negative values for  $c$  correspond to liberal observers and imply that misses are disproportionately reduced at the cost of higher false alarm rates (Fig. 1). Conversely, positive values for  $c$  indicate a conservative response bias that reduces false alarms at the cost of more misses (Fig. 1). Further comprehensive reviews on signal detection and measures of sensitivity and response bias are offered by Green & Swets (1966), Macmillan & Creelman (1990), McNicol (1972) and Monk & Eiser (1980).

## The experiment

To assess performances in the identification of turbate micro-structures and to determine the extent to which this identification performance is affected by experience, we asked people with various levels of experience to classify photomicrographs of glacial diamictons as either containing turbate micro-structures (signal plus noise) or not (noise alone).

### Participants

Fifty-two participants were recruited through word of mouth. Participants were asked to rate their own experience levels, differentiating between microscopic work in general and the analysis of turbate micro-structures in particular. Relative four-point scales were used, with '1' indicating 'no experience' and '4' indicating 'a high level of experience'. We added up the scores from these ratings to obtain an overall measure of experience. Participants were subsequently grouped as follows: those with 'No experience' (complete novices, 2 points, 12 participants), 'Beginners' (3–4 points, 20 participants), those with 'Moderate' levels of experience (5–6 points, 12 participants), and those with high levels of experience (referred to as the 'Expert' group hereafter; 7–8 points, 8 participants).

### Trials

The selection of trials for the experiment occurred following a process in which five people with various levels of experience (and who did not participate further in the experiment) were asked to establish for ~80 photomicrographs whether they contained a signal or not. Only those trials for which there was unanimous agreement among the five observers were included in the study. This procedure resulted in a set of 45 trials – 21 with a turbate micro-structure (noise plus signal) and 24 without a turbate micro-structure (noise alone, no signal). The 45 photomicrographs were presented in a standard rectangular format (same aspect ratio for all). Figure 2 shows examples for both categories.

### Procedure

All participants, regardless of experience, received detailed instructions on the identification of turbate micro-structures. Instructions to the participants comprised a visual (see Fig. 3) as well as a written explanation. The latter was based on a definition of a turbate micro-structure following van der Meer (1993): 'Turbate micro-structures can be found in thin sections of sediments. They can take the form of either a circular (elliptical) arrangement of "satellite" grains around a larger core grain or clast, or they consist of a circular arrangement of grains without a larger core, as shown below. The size of turbate features may vary

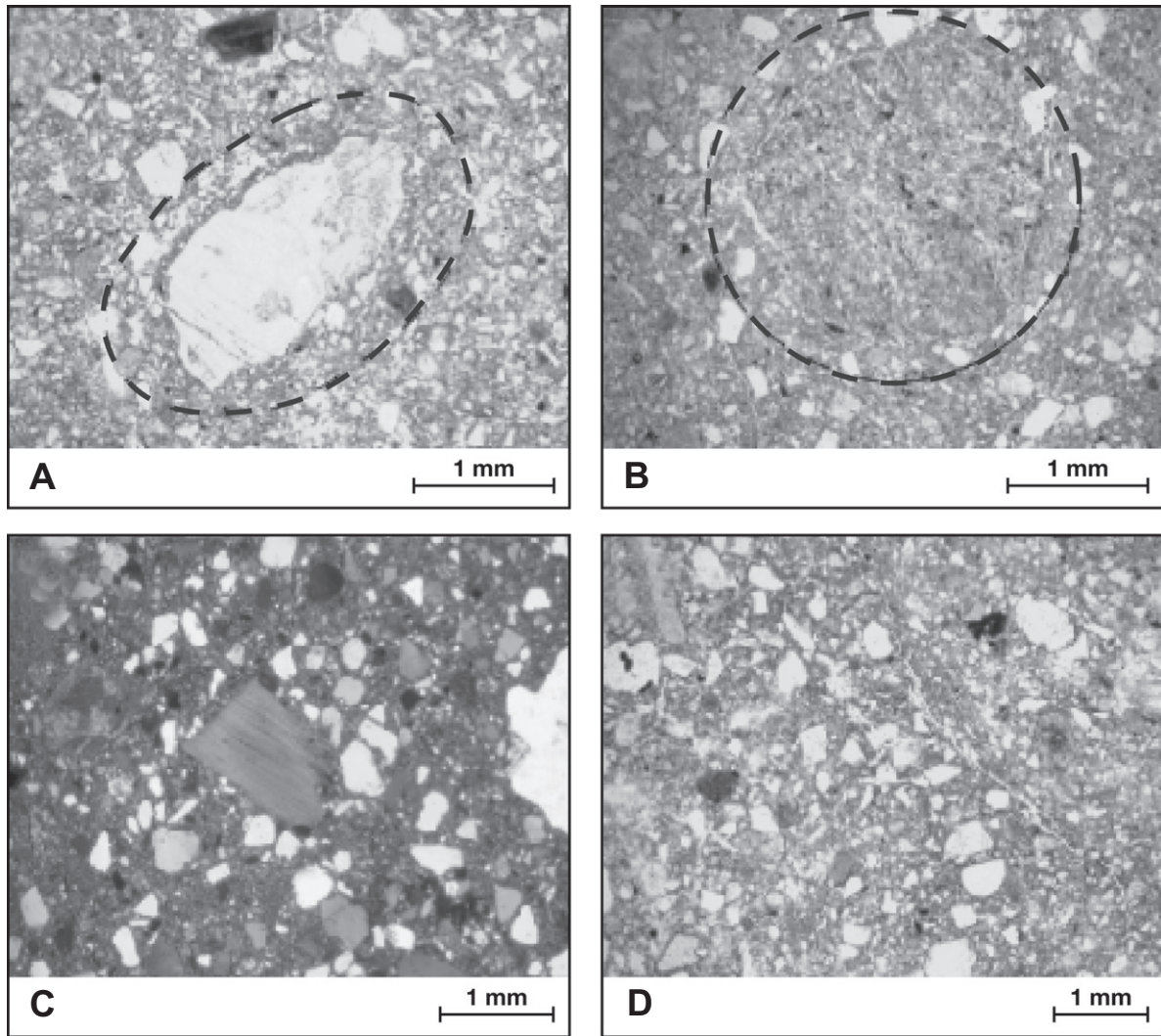


Fig. 2. Examples of photomicrographs used in the experiment. Turbate micro-structures in images A and B are indicated (trials with signal and noise – annotations not in original images). Images C and D are examples of trials without signal (just noise, no turbate micro-structure).

significantly, while the number of satellite grains is irrelevant. The only requirement is that long axes of most satellite grains describe a circular or elliptical pattern and/or show positions parallel to the surface of the core.'

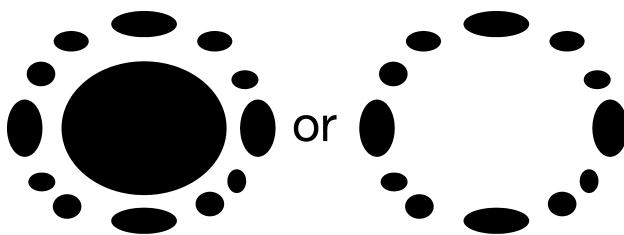


Fig. 3. Instructional image: examples of turbate micro-structure with 'core' (left) and turbate micro-structure without 'core' (right).

Following the instruction, photomicrographs were presented one at a time, and participants were asked to indicate on a form whether or not the respective photomicrograph contained a turbate micro-structure before the next photomicrograph was shown. The order of photomicrographs was determined by random number generation (Haahr 1998) and was kept the same for all participants.

### Results

Figure 4A shows sensitivity and bias data for individual participants. Figure 4B shows the envelopes encompassing clouds of data points per experience group. Also indicated in Fig. 4B are the means per experience group (calculated using absolute values of  $c$ ; see Table 1). In the beginners group, two participants

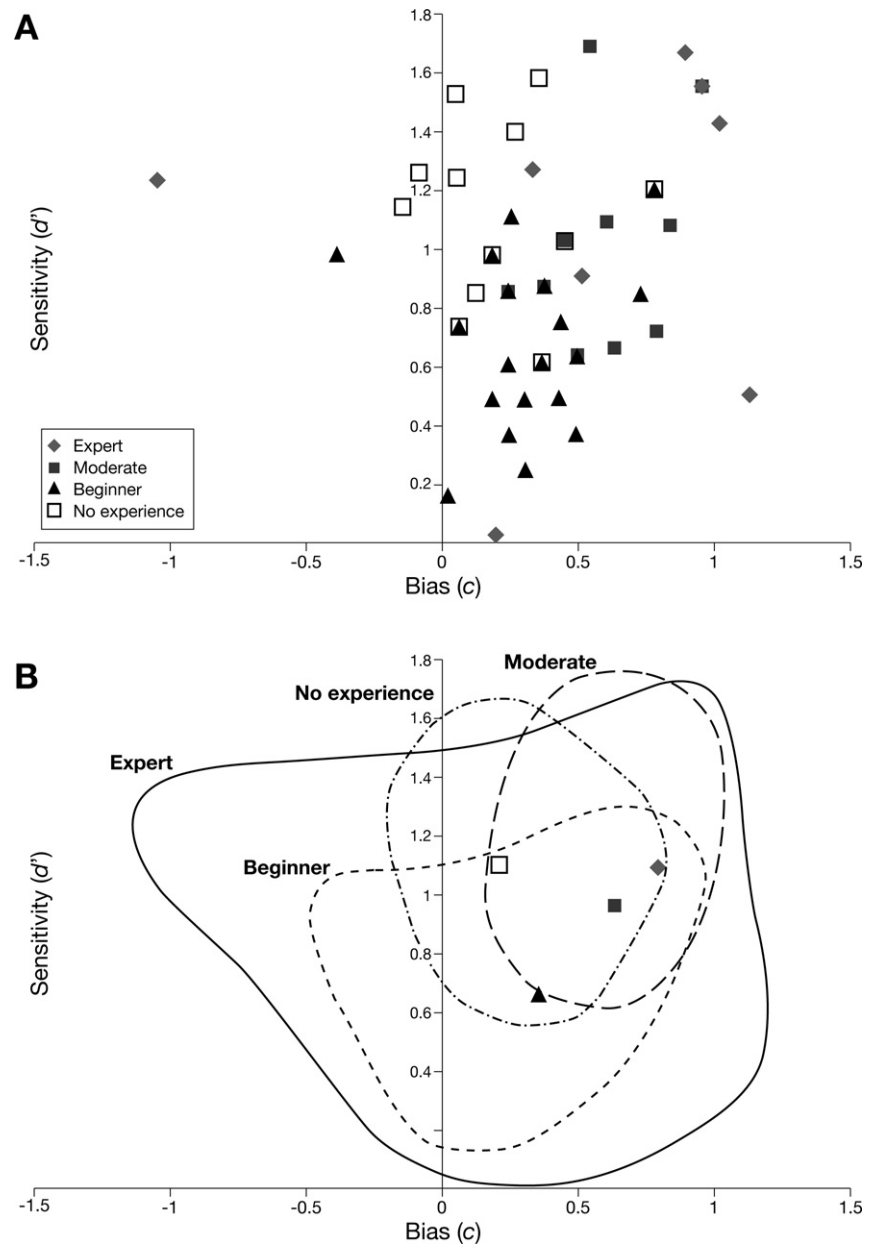


Fig. 4. Bias  $c$  versus sensitivity  $d'$ . A. All data. B. Envelopes highlight the ranges within experience groups. Also indicated are the absolute means. For further explanation see text.

scored negative  $d'$ -values. Given that chance performance (i.e. pure guessing) would correspond to a  $d'$ -value of 0, a negative value may be expected to represent either a capable participant's deliberate attempt to perform poorly, or a participant's misunderstanding

Table 1. Mean and standard deviation: bias, sensitivity and absolute bias per experience group (mean and standard deviation).

Group	$d'$	$c$	$ c $
Expert	$1.076 \pm 0.563$	$0.498 \pm 0.712$	$0.761 \pm 0.360$
Moderate	$0.986 \pm 0.342$	$0.574 \pm 0.204$	$0.574 \pm 0.204$
Beginner	$0.646 \pm 0.267$	$0.277 \pm 0.234$	$0.320 \pm 0.166$
No experience	$1.130 \pm 0.301$	$0.207 \pm 0.257$	$0.244 \pm 0.218$

of instructions. In any case it indicates a lack of compliance with the instructions, which is why the data from these two participants were excluded from the analyses.

The second thing to note is that in the Expert group, one participant showed a particularly low (negative)  $c$ -value, that is, a very liberal response bias. This is an interesting result that will be discussed further below.

To assess whether the respective experience groups differed statistically in sensitivity  $d'$  and bias  $c$ , separate analyses of variance (ANOVAs) were conducted for these two dependent variables (we adopted the common significance level of  $\alpha=0.05$  for all tests). It was

found that sensitivity and bias differed significantly across levels of expertise ( $F[3, 46]=5.709$ ,  $MSE=0.709$ ,  $p=0.002$  and  $F[3, 46]=2.997$ ,  $MSE=0.369$ ,  $p=0.040$ , respectively). By investigating absolute bias, which also differed significantly between groups ( $F[3, 46]=11.434$ ,  $MSE=0.586$ ,  $p<0.001$ ), it was confirmed that the observed differences in 'raw' bias were not solely the result of differences in the bias 'direction'. Sensitivity and absolute bias  $|c|$  also differed significantly between groups that had any experience (i.e. when the No experience group was excluded;  $F[2, 35]=5.130$ ,  $MSE=0.691$ ,  $p=0.011$  and  $F[2, 35]=11.373$ ,  $MSE=0.596$ ,  $p<0.001$ , respectively), and the difference in raw bias between groups with any experience fell just short of statistical significance ( $F[2, 35]=2.487$ ,  $MSE=0.351$ ,  $p=0.098$ ).

Table 1 and Fig. 4 reveal that sensitivity was nominally higher for the No experience group than for the Experts, but this difference was not found to be significant ( $t[18]=-0.280$ ,  $SE=0.193$ ,  $p=0.783$ ; all  $p$ -values for  $t$ -tests reported here are two-sided without corrections for multiple comparisons). Significant differences in sensitivity between pairs of groups were found to exist between the Experts and the Beginners ( $t[24]=2.680$ ,  $SE=0.161$ ,  $p=0.013$ ), between the Moderately experienced and the Beginners ( $t[28]=3.063$ ,  $SE=0.111$ ,  $p=0.005$ ), as well as between the Beginners and the No experience group ( $t[24]=-4.624$ ,  $SE=0.105$ ,  $p<0.001$ ). Raw bias only differed significantly between the Moderately experienced group and the two less experienced groups ( $t[22]=3.882$ ,  $SE=0.095$ ,  $p=0.001$  and  $t[28]=3.579$ ,  $SE=0.083$ ,  $p=0.001$  for the comparisons with the No experience group and the Beginners, respectively), whereas absolute bias differed significantly between the Experts and the No experience group ( $t[18]=4.109$ ,  $SE=0.222$ ,  $p<0.001$ ), the Experts and the Beginners ( $t[24]=4.332$ ,  $SE=0.183$ ,  $p<0.001$ ), as well as between the Moderately experienced group and the Beginners ( $t[28]=3.746$ ,  $SE=0.083$ ,  $p<0.001$ ).

## Discussion

Figure 4B and Table 1 suggest a trend in sensitivity-to-signal  $d'$  with increasing experience (Beginner – Moderate – Expert), as may be expected. This trend is confirmed to be statistically significant. However, there is also an apparent and statistically corroborated increase of absolute bias (i.e. deviations from neutral) with experience, with particularly high inter-individual variability among the Experts.

Very interesting to note are the sensitivity and bias data for the complete novices (the No experience group). While the mean bias is nominally the lowest of all groups, the mean sensitivity value is nominally even higher than that of the Experts. Although one Expert showed an exceptionally low sensitivity (see Fig. 4A),

which obviously pulled down the mean of the entire group (Fig. 4B), the excellent performance of the complete novices is still noteworthy, particularly when compared with that of the Beginners and the group of Moderately experienced analysts. We postulate that the high sensitivity for the group without relevant experience is due to members of this group frequently referring back to the instructions – a strategy that may have mitigated and compensated for the lack of experience. Informal questioning of participants after the experiment suggested that participants from this group indeed relied more heavily on the instructions during the course of the experiment. Following this reasoning, participants with more experience may have overestimated their abilities and not have 'relied on' the instructions to the same extent.

Participants showed a tendency to be conservative in their detection judgments. Across all experience levels, false alarm rates were in most cases significantly lower than miss (i.e. 1-hit) rates, which translated to positive values of bias, almost across the board. This overall conservative response bias might reflect participants' evaluation of consequences of different types of error. We speculate that participants generally err on the side of caution, because they perceive that erroneously detecting a non-existent signal is worse than failing to detect a signal that is actually there. Only four participants (see Fig. 4A) showed a negative (i.e. liberal) response bias. We did not find response criteria to be 'more neutral' for Experts than for participants with less experience. In fact, all participants with a near-neutral response bias were from either the No experience or the Beginner group.

Participants from the Expert group exhibited both the most conservative and the most liberal response biases. This suggests that, in addition to the aforementioned 'cautious' response, which seems to apply irrespective of experience level, there are other factors that play a more prominent role with increasing experience. We propose that the heterogeneity of the learning history is one of these. Russo *et al.* (2008) demonstrated that a drive for consistency can be a primary forcing in the development of a preference or bias. Similarly, it has been shown by Holyoak & Simon (1999) that processes of subconscious human reasoning can lead to a positive reinforcement of an initial premise.

Another likely (and associated) source of variability in bias within the Expert group is the learned base-rate of thin-section samples containing turbate microstructures. A Bayesian ideal observer making decisions about noisy stimuli (such as the photomicrographs used in our experiment) may be expected to integrate prior knowledge about the likelihood of a signal with new observations. It is possible, therefore, that Experts with different learned base-rates (and/or different inclinations to employ this prior knowledge) develop

different 'preferences', which are in turn manifested in different response biases. In other words, the variability in response bias observed among the Experts may reflect adaptive decision processes that serve to reduce errors in settings with signal base-rates that match previously learned contingencies.

## Conclusions

This study highlights the role that experience may play in the interpretation of thin sections of (glacial) sediments. The results of our experiment suggest that experience increases an observer's sensitivity or ability to detect structures (here: turbate micro-structures). The experiment also demonstrates that expert levels of sensitivity can be obtained by complete novices, if they are provided with detailed instructions to which they can refer back during the identification process.

The results of the experiment suggest that decision-making can be controlled by psychological factors, including cautiousness – which would explain the observed overall conservative response bias – and the drive for consistency. The results also suggest that the response bias may be amplified by previous experiences such as the learned base-rate of samples containing signals. Dependent upon how such experiences are projected onto new detection tasks, and how an observer's preference is developed, this may well lead to a distinctly conservative or liberal (i.e. non-neutral) response bias.

Whereas the outcome of our study is encouraging in that it confirms the prominent role of experience (and training) in improving detection performance, it also carries an implicit warning in that several psychological factors can affect an analyst's perception of structures and detection judgments. Obviously, in the context of glacial micromorphology, a non-neutral response bias would lead to a higher number of erroneous interpretations of sediments than would be necessary based on the perceptual limits. We recommend therefore that, where possible, sediments are investigated by more than one analyst, and also that, rather than single micro-structures, a set of multiple micro-structures should serve as the basis for any interpretation.

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